FELT JOB INSECURITY AND UNION MEMBERSHIP: THE CASE OF TEMPORARY WORKERS

Nele DE CUYPER, Hans DE WITTE, Magnus SVERKE, Johnny HELGREN, Katharina NASWALL

Research Group Work, Organization and Personnel Psychology, KU Leuven, Belgium
North-West University, South Africa
Department of Psychology, Stockholm University, Sweden
University of Canterbury, New Zealand

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The present study investigates the relationship between felt job insecurity and union membership accounting for potential differences between temporary and permanent workers. Consistent with the idea that felt job insecurity leads workers to seek social protection from the unions, and with earlier studies, we hypothesize a positive relationship between felt job insecurity and union membership (Hypothesis 1). Furthermore, we argue that this relationship may be stronger among temporary compared with permanent workers (Hypothesis 2): insecure temporary workers are in a situation of ‘double vulnerability’, hence they have strong motives for unionization. Hypotheses are tested in a cross-sectional sample of 560 Flemish (Dutch-speaking part of Belgium) workers. Our results were as follows: the relationship between felt job insecurity and union membership was not significant. The interaction term between contract type and felt job insecurity was significantly related to union membership: the relationship between felt job insecurity and union membership was positive among temporary workers, but not among permanent workers. This pattern of results may inspire unions to target future recruitment strategies on temporary workers. A route for future research could be to test our hypotheses also longitudinally.

Keywords: contingent employment, fixed-term employment, job insecurity, social protection, union

E-mail: nele.decuyper@ppw.kuleuven.be
INTRODUCTION
Antecedents of union membership have been sought at different levels: economic (e.g., unemployment rate), institutional (e.g., union-affiliated unemployment benefits), personal (e.g., age, occupational position) and psychological (Stinglhamber, Gillis, Teixeira, & Demoulin, 2013). Psychological studies have mostly taken the perspective that unionization is a way to voice dissatisfaction with existing working conditions (Bamberger, Kluger, & Suchard, 1999; Buttigieg, Deery, & Iverson, 2007; Johnson & Jarley, 2004). A particular cause for dissatisfaction in today’s work environment and one that has traditionally been considered a concern promoting union membership is felt job insecurity (Sverke & Hellgren, 2001; Waddington & Whitston, 1997). Felt job insecurity concerns the employee’s perception and concern about potential involuntary job loss (De Witte, 1999, 2005; Sverke, Hellgren, & Näsvall, 2002; Vander Elst, De Witte, & De Cuyper, 2013). Workers who fear they might lose their job may turn to the union to seek social support and protection.

Previous studies have established a positive association between felt job insecurity and union membership (e.g., Allvin & Sverke, 2000; Bender & Sloane, 1999; Nätti, Happonen, Kinnunen, & Mauno, 2005; Sverke et al., 2004). We will take these studies one step further by accounting for type of contract, temporary versus permanent employment in particular. Temporary employment refers to ‘dependent employment of limited duration’ (OECD, 2002, p. 170), for example in the form of fixed-term employment (for an overview, see De Cuyper et al., 2008). We believe the issue of temporary employment in union research requires more specific attention as temporary workers have strong motives for union membership: temporary workers are generally more insecure than permanent workers (De Cuyper & De Witte, 2006, 2007), and they may lack other forms of social support at work as they relatively often change employers. Accordingly, a plausible assumption is that the relationship between felt job insecurity and union membership is particularly strong among temporary compared to permanent workers. Surprisingly, temporary employment has not yet attracted much attention in the realm of union research: earlier studies have mostly focussed upon permanent workers (Goslinga & Sverke, 2003), thereby largely excluding temporary workers from the analyses (Bender & Sloane, 1999).

In response, it is the intent of this study to investigate the relationship between felt job insecurity and union membership among temporary and permanent workers from two sectors in Flanders, Belgium. Results are based on cross-sectional data. From a practitioners’ point of view, our findings may
help to understand the challenges the union is facing: dealing with increasing levels of felt job insecurity against the background of a growing number of temporary workers.

**JOB INSECURITY AND UNION MEMBERSHIP**

The dominant view is that job insecurity may elicit and strengthen union membership. The main argument is that potential and current union members seek social protection. Social protection should be understood broadly in terms of preventing and managing situations that may adversely impact the workers’ well-being (United Nations Research Institute for Social Development, 2010), implying also protection against job loss. In this respect, the protection of workers against job loss is central to employees’ decision to unionize and to stay unionized (Barling, Fullager, & Kelloway, 1992; Gallagher & Strauss, 1991; Stingelhämmer et al., 2013; Sverke & Hellgren, 2001; Waddington & Whitston, 1997).

This view aligns with the frustration-aggression thesis that explains union membership as a function of employees’ frustration with dissatisfying working conditions (Bryson, Cappellari, & Lucifora, 2004; Hammer & Avgar, 2005; Nätti et al., 2005; Sverke et al., 2004) and their need to voice their frustration (Sverke & Hellgren, 2001).

Furthermore, it finds support in earlier studies. Direct evidence comes from studies that have established a positive relationship between felt job insecurity and union membership (e.g., Allvin & Sverke, 2000; Bender & Sloane, 1999; Nätti et al., 2005; Sverke et al., 2004), though some exceptions exist (Hartley, Jacobson, Klandermans, & Van Vuuren, 1991). Indirect evidence comes from the observation that felt job insecurity feeds pro-union attitudes (Sverke & Goslinga, 2003; Sverke et al., 2004) and reduces turnover from the union (Sverke & Goslinga, 2003; Sverke et al., 2004). In line with the dominant view and evidence to date, our hypothesis is as follows:

H1: Felt job insecurity relates positively to union membership.

**THE CASE OF TEMPORARY WORKERS**

Most research has concerned the situation of permanent workers, while the relationship between felt job insecurity and union membership may be different for temporary workers. Arguments for the differential relationships are plenty, and they fall in two categories, depending on whether the relationship is assumed weaker of stronger.

To begin with, felt job insecurity may not be a cause for unionization among temporary workers to the same extent as
it is among permanent workers. One reason is that temporary workers may fear that unionization reduces future prospects in the organization based on the idea that employers are more likely to hire non-unionized workers or may discriminate union members when hiring new personnel (for a discussion, see e.g., Cooke, 1985). Many temporary workers see their assignment as a stepping stone to permanent employment (De Cuyper & De Witte, 2008), and they may more easily gain a permanent job through individualistic actions that please the employer, for example, excellent performance, citizenship and impression management. Another reason is that temporary workers, unlike permanent workers, may not see job insecurity as a breach of the psychological contract they have with the employer. Instead, they see job insecurity as part of the deal (De Cuyper & De Witte, 2006, 2007; Klandermans, Klein Hesselink, & Van Vuuren, 2010). Accordingly, temporary workers may not feel the need to turn to the union in the face of job insecurity. In line with this reasoning, previous studies have demonstrated that felt job insecurity does not relate (as strongly) to strain among temporary workers (compared with permanent workers) (Bernhard-Oettel, Sverke, & De Witte, 2005; De Cuyper & De Witte, 2006, 2007; De Witte & Näswall, 2003; Klandermans et al., 2010; Mauno, Kinnunen, Mäkikangas, & Nätti, 2005). Hence, frustration with job insecurity may not be the driving force for unionization among temporary workers, while it is among permanent workers.

Alternatively, felt job insecurity may relate more strongly to unionization among temporary compared with permanent workers. First, temporary workers are in a notoriously weaker position to voice their insecurity concerns: they may not know the organizational procedures to address their concerns, they may not have sufficiently strong professional networks in the organization and feel isolated, or they may be afraid of possible reprimands. Accordingly, they may see the union as a way to voice their opinions and strengthen their position though in a more anonymous way. This aligns with the frustration-aggression thesis (Bryson et al., 2004; Hammer & Avgar, 2005; Nätti et al., 2005; Sverke et al., 2004): union membership is the result of individual’s frustration, dissatisfaction or alienation at work. Second, Belgian unions are involved in the administration of unemployment insurances, which may be appealing particularly for job insecure temporary workers: they face ‘double’ vulnerability (i.e., insecurity and temporarily employed) and may want to anticipate job loss.
We follow this last line for two reasons. First, it ties in with our earlier argument that seeking social protection is an important motive for unionization. Seeking social protection is particularly relevant for temporary workers: they do not have a broad social network within the organization. Second, empirical evidence seems to support the idea of a stronger relationship between felt job insecurity and union membership among temporary compared with permanent workers. Nätti et al. (2005) reported a similar interaction effect based on the Finnish Working Conditions Survey for the year 2003. Indirect evidence comes from the study by Furaker and Berglund (2003): temporary workers compared with permanent workers were somewhat more likely to agree that unions are needed. In line with this evidence, our second hypothesis is as follows:

H2: The positive relationship between felt job insecurity and union membership is stronger among temporary compared with permanent workers.

METHOD

Data Collection

Data was collected in divisions of seven Belgian companies within Flanders (N = 567) that agreed to participate in a survey on the quality of working life (see Guest, Isaksson, & De Witte, 2010). Two sectors were selected for this study, namely industry with one large company (N = 263) and retail with six smaller organizations (N = 304). The response rate for the industrial setting was 87.6%: this organization was highly committed to the research and invested heavily in communication towards employees (e.g., by distributing flyers). Response rates for the retail sector varied between 33% and 58% in five out of six organizations. One retail organization did not have a HR department to coordinate data collection, which may explain the lower response rate of 20.8%. Sectors and organizations were recruited for three reasons: possibilities for generalizing findings, number of temporary workers and variation in policies towards temporary employment. Potential respondents were asked to fill out the questionnaire during working time, either during group sessions facilitated by the researchers or individually, or, if preferred, at home. Confidentiality and voluntary participation was stressed, and feedback at the company level was guaranteed.

Agency workers were excluded from the analyses owing to their specific triangular employment relationship and to specific regulations for unionization in the agency sector. More-
over, the small number of agency workers in our sample did not allow more detailed analyses. This led to a total sample size of 560 respondents.

**Respondents**

About one respondent out of three (N = 189; 33.7%) was temporarily employed on a fixed-term contract. The other respondents were permanently employed on an open-ended contract (N = 371; 66.3%). This was not representative for the Belgian labour market which has less than ten per cent temporary workers (De Cuyper, De Witte, & Isaksson, 2005), but instead was the result of the researchers’ effort to sample temporary workers. About two respondents out of three were union members (N = 188; 66.3%), which is slightly higher than the population percentage (60%; Sverke et al., 2004).

More females (N = 358; 65.3%) than males (N = 190; 34.7%) participated in the study. The majority of the respondents did not follow education beyond high school (N = 422; 75.4%), which related to the relatively high share of blue-collar workers (N = 351; 65.2%) compared with white-collar workers (N = 187; 34.8%) in this study. Mean age of the sample was 35 years (SD = 10), ranging from 18 to 58 years. Mean tenure was 10 years (SD = 9), with a variation from 1 month to 39 years. Respondents worked on average 31 hours per week (SD = 9), and roughly half of the respondents worked part-time (47.7%).

There were differences between the temporary and the permanent sample. First, temporary workers were more likely to be employed in the retail sector (64.6%) than in the industrial sector (35.4%). This difference was smaller for permanent workers: 48.2% versus 51.8%, respectively, $\chi^2(1, N = 560) = 13.39, p < 0.001$. Second, unionization rate was lower among temporary workers (53.7%) than among permanent workers (72.7%), $\chi^2(1, N = 558) = 20.10, p < 0.001$. Third, the temporary sample included relatively more women (74.3%) than the permanent sample (60.8%), $\chi^2(1, N = 548) = 9.80, p < 0.01$. Fourth, more permanent workers (80.9%) than temporary workers (64.6%) followed higher education, $\chi^2(1, N = 560) = 17.94, p < 0.001$. Furthermore, temporary workers (M = 29 years; SD = 9.92) were on average younger than permanent workers (M = 37 year; SD = 9.01), t(533) = -9.22, p < 0.001, and they had lower organizational tenure (M_{temporary} = 3 year; SD = 4.53; M_{permanent} = 14 year; SD = 9.10), t(552) = -15.16, p < 0.001. Finally, temporary workers worked fewer hours per week (M = 27 hours/week; SD = 12.01) than permanent workers (M = 33 hours/week; SD = 7.64), t(547) = -6.48, p < 0.001. No such differences between the temporary and the permanent sample were found for occupational position (blue-collar versus white-collar workers), $\chi^2(1, N = 538) = 0.42, p = 0.57$. 

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Measures

In all analyses, we controlled for the following variables: organizational tenure (years), average weekly working hours, gender (0 = female; 1 = male) and occupational position (0 = blue-collar workers; 1 = white-collar workers). We selected these controls because they are important in relation to union membership (Monnot, Wagner, & Beehr, 2011) and felt job insecurity (Näswall & De Witte, 2003). We did not control for age because of its high correlation with tenure ($r = 0.70$, $p < 0.01$). Similarly, we did not control for sector because of its high association with occupational position, $\chi^2(1, N = 538) = 233.73$, $p < 0.001$: the majority of blue-collar workers (70.7%) worked in industry, and the large majority of white-collar workers (98.4%) in retail. Note, however, that we performed the analyses with different subsets of control variables (e.g., age and sector instead of tenure and occupational position), with essentially the same results.

Union membership was coded 0 for non-members and 1 for members.

Contract type was coded 0 for temporary workers and 1 for permanent workers.

Felt job insecurity was assessed using four items developed by De Witte (2000) and validated by Vander Elst et al. (2013). Sample items were ‘I feel insecure about the future of my job’ and ‘I think I might lose my job in the near future’. Reliability (Cronbach’s alpha) was 0.85.

Information about means, standard deviations and correlations is to be found in Table 1.

<table>
<thead>
<tr>
<th>TABLE 1</th>
<th>Means, standard deviations and correlations between the study variables</th>
</tr>
</thead>
<tbody>
<tr>
<td>M</td>
<td>SD</td>
</tr>
<tr>
<td>1. Union membership</td>
<td>0.19**</td>
</tr>
<tr>
<td>2. Contract type (permanent)</td>
<td>2.52</td>
</tr>
<tr>
<td>3. Job insecurity</td>
<td>9.99</td>
</tr>
<tr>
<td>4. Tenure</td>
<td>31</td>
</tr>
<tr>
<td>5. Working hours</td>
<td>0.07</td>
</tr>
<tr>
<td>6. Gender (male)</td>
<td>-0.08</td>
</tr>
<tr>
<td>7. Occupational position (white-collar)</td>
<td>-0.08</td>
</tr>
</tbody>
</table>

Note: Union membership: 0 = non-member, 1 = member; Contract type: 0 = temporary contract, 1 = permanent contract; Gender: 0 = female, 1 = male; Occupational position: 0 = blue-collar worker, 1 = white-collar worker. * $p < 0.05$; ** $p < 0.01$

Analyses

We used hierarchical moderator logistic regression to test our hypotheses. The control variables were entered at the first step, contract type in the second step, and felt job insecurity in the
third step. Finally, we included the interaction term between contract type and felt job insecurity. In doing so, we centred the predictor variables before creating the interaction term. Upon significance, the interaction was plotted.

**RESULTS**

Table 2 summarizes the results of the hierarchical moderator logistic regression. From the control variables, only weekly working hours was associated with union membership: with each unit increase in weekly working hours, the odds of union membership go up by a multiplicative factor of 1.05 (Table 2, Step 1). Furthermore, permanent compared with temporary workers were about twice as likely to report union membership (Table 2, Step 2).

<table>
<thead>
<tr>
<th></th>
<th>Step 1</th>
<th>Step 2</th>
<th>Step 3</th>
<th>Step 4</th>
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<tr>
<td></td>
<td>β</td>
<td>Exp(β)</td>
<td>β</td>
<td>Exp(β)</td>
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<tr>
<td>Tenure</td>
<td>0.01</td>
<td>1.01</td>
<td>-0.01</td>
<td>0.99</td>
</tr>
<tr>
<td>Working hours</td>
<td>0.05***</td>
<td>1.05</td>
<td>0.05***</td>
<td>1.05</td>
</tr>
<tr>
<td>Gender (male)</td>
<td>0.03</td>
<td>1.03</td>
<td>-0.03</td>
<td>0.97</td>
</tr>
<tr>
<td>Occupational position (white-collar)</td>
<td>-0.10</td>
<td>0.91</td>
<td>-0.14</td>
<td>0.87</td>
</tr>
<tr>
<td>Contract type (permanent)</td>
<td>0.76***</td>
<td>2.15</td>
<td>0.80***</td>
<td>2.23</td>
</tr>
<tr>
<td>Job insecurity</td>
<td>0.04</td>
<td>1.04</td>
<td>0.28</td>
<td>1.32</td>
</tr>
<tr>
<td>Contract type x job insecurity</td>
<td>-0.46*</td>
<td>0.63</td>
<td></td>
<td></td>
</tr>
<tr>
<td>χ²</td>
<td>32.76</td>
<td>9.37</td>
<td>0.13</td>
<td>4.57</td>
</tr>
<tr>
<td>(4, N = 560)</td>
<td>(1, N = 560)</td>
<td>(1, N = 560)</td>
<td>(1, N = 560)</td>
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<tr>
<td>p &lt; 0.001</td>
<td>p &lt; 0.01</td>
<td>p = 0.72</td>
<td>p &lt; 0.05</td>
<td></td>
</tr>
<tr>
<td>R² Nagelkerke</td>
<td>0.09</td>
<td>0.11</td>
<td>0.11</td>
<td>0.12</td>
</tr>
</tbody>
</table>

Note: Union membership: 0 = non-member, 1 = member; Contract type: 0 = temporary contract, 1 = permanent contract; Gender: 0 = female, 1 = male; Occupational position: 0 = blue-collar worker, 1 = white-collar worker. Note: β = log odds; Exp(β) = odds.

Our first hypothesis concerned the association between felt job insecurity and union membership. However, felt job insecurity was not significantly related to union membership, and hence Hypothesis 1 was not supported (Table 2, Step 3).

Hypothesis 2 concerned the interaction between contract type and felt job insecurity in relation to union membership. The interaction term contributed to explaining union membership (Table 2, Step 4). Figure 1 shows that felt job insecurity was positively related to union membership among temporary workers, but not among permanent workers. This supported Hypothesis 2.
Concerns about job insecurity are central to the union's discourse, and scholars have identified felt job insecurity as a motive for joining as well as staying with the union (e.g., Allvin & Sverke, 2000; Bender & Sloane, 1999; De Witte et al., 2008; Nätti et al., 2005; Sverke et al., 2004): current and potential union members seek social protection when they anticipate job loss. Accordingly, we hypothesized a positive relationship between felt job insecurity and union membership (Hypothesis 1), but we did not establish this relationship in our sample.

One explanation could be that most workers joined the union early in their career and for reasons other than felt job insecurity: dissatisfaction is generally not the main cause for union membership and participation (Klandermans, 1986). Workers join the union because there is a strong tradition of unionization, for example in Belgium, or because they have instrumental reasons such as representation in collective bargaining or access to extra services such as free legal support in case of disputes and social assistance (Stinglhamber et al., 2013).

While this explanation may apply to permanent workers, unionization at labour market entry is less likely among temporary workers: the union's recruitment efforts are not typically focused upon temporary workers, and temporary workers may not hear about positive union experiences from their colleagues because their networks in the organization are generally weaker (Visser, 1995). This may explain why permanent workers are about twice as likely to be union members (Goslinga & Klandermans, 2001; Nätti et al., 2005), also in countries with high union density. The implication could be that temporary workers are more actively seeking the union's support on their own initiative, and this may be triggered by
felt job insecurity: temporary workers who feel insecure may hope to find some support to address their precarious situation and they may seek information about unemployment insurance, or anticipate the help unions provide once they are unemployed. Accordingly, a plausible hypothesis is that felt job insecurity relates to union membership among temporary workers but less so among permanent workers (Hypothesis 2), which found support in the study by Nätti et al. (2005) and in the present study: felt job insecurity was positively related to union membership among temporary workers, but not among permanent workers.

The message to unions is that temporary workers are a valuable group to consider for prospective union membership: unionization among temporary workers is relatively low, while at the same time reasons to unionize among temporary workers are strong. Temporary workers may however experience difficulties in finding their way to the union: Belgian unions are represented at the workplace as a means to develop and enhance union commitment through daily contact with union representatives. Temporary workers frequently change jobs or cycle between spells of short-term employment and unemployment, and they are less likely to be integrated in professional networks that favour unionization or share union experiences (Nätti et al., 2005). This presents challenges in the recruitment, organization and representation of temporary workers (Allvin & Sverke, 2000; Goslinga & Sverke, 2003; Nätti et al., 2005; Sverke et al., 2004). Furthermore, the voting and election systems tend to exclude temporary workers (Kerkhof, Winder, & Klandermans, 2005). The implication is that unions may be unaware of the specific needs and concerns of temporary workers, or they may see few incentives to represent temporary workers in collective bargaining.

Limitations

The results of the present study should be interpreted with a degree of caution and accounting for the following three limitations. First, we used cross-sectional survey data, which obviously limits causal interpretations. We assumed that felt job insecurity ‘leads to’ union membership, but it could be argued that union membership may also feed feelings of job insecurity: union membership may raise awareness about potential risks related to job insecurity (Sverke et al., 2004). A more provocative idea concerns potential unintended side effects of unionization (Bender & Sloane, 1999): unions may reduce employment opportunities in the organization by negotiating higher wages, which then may lead to an overall increase in feelings of job insecurity. However, union membership is un-
likely to be a cause of felt job insecurity among temporary workers: felt job insecurity is inherent to temporary employment. Furthermore, our hypothesis was based on earlier studies on the relationship between felt job insecurity and union membership (Allvin & Sverke, 2000; Bender & Sloane, 1999; De Witte et al., 2008; Nätti et al., 2005; Sverke et al., 2004).

Second, some caution is warranted when the aim is to generalize our findings to the Belgian/Flemish population or to other countries. Our sample was not representative for the Belgian/Flemish population. This was related to our strategy of oversampling temporary workers. Moreover, we focused on the dominant group of fixed-term contract workers in Belgium, and hence we did not account for heterogeneity in temporary contracts. We realize that results may be different among other groups of temporary workers, agency workers or on-call workers, for example. Similarly, results may be conditional upon the unions’ recruitment strategies, voting and election systems and organization at the workplace. In this respect, we realize that our arguments may be bound to the Belgian context, with a relatively strong union movement. This is a common problem in the realm of union research. For example, the European and US union literature show some differences based on how unions are organised.

Third, the explained variance was about 12%. This suggests that there are other and potentially more important predictors of union membership. In this respect, Sverke et al. (2004) have highlighted attitudes towards the union (e.g., union satisfaction, union commitment), the job (e.g., job satisfaction) and the organization (e.g., affective organizational commitment). Future research may want to investigate whether these factors are equally important to temporary compared with permanent workers.

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Doživljaj nesigurnosti posla i članstvo u sindikatu: slučaj privremenih radnika

Nele DE CUYPER\textsuperscript{a}, Hans DE WITTE\textsuperscript{a,b}, Magnus SVERKE\textsuperscript{b,c}, Johnny HELLGREN\textsuperscript{c}, Katharina NASWALL\textsuperscript{c,d}
\textsuperscript{a} KU Leuven, Belgija
\textsuperscript{b} Sveučilište North-West, Južna Afrika
\textsuperscript{c} Sveučilište u Stockholmu, Švedska
\textsuperscript{d} Sveučilište Canterbury, Novi Zeland

Ova studija istražuje odnos između doživljaja nesigurnosti posla i članstva u sindikatu, objašnjavajući potencijalne razlike između privremeno zaposlenih radnika i onih u stalnom radnom odnosu. U skladu s tezom da doživljaj nesigurnosti posla navodi radnike da socijalnu zaštitu zatraže od sindikata, a i u skladu s ranijim istraživanjima, postavlja se hipoteza pozitivne povezanosti između doživljaja nesigurnosti posla i članstva u sindikatu (Hipoteza 1). Nadalje, tvrdi se da ova povezanost može biti jača kod privremenih u usporedbi sa stalnim radnicima (Hipoteza 2): nesigurni privremeni radnici nalaze se u položaju "dvostruke ranjivosti", stoga imaju snažne motive za sindikalno organiziranje. Hipoteze su testirane na uzorku poprečnoga presjeka, koji se sastojao od 560 flamanskih radnika (nizozemsko govorno područje u Belgiji). Rezultati su ovakvi: povezanost doživljaja nesigurnosti posla i članstva u sindikatu nije bila značajna. Interakcija između vrste ugovora i doživljaja nesigurnosti posla bila je značajno povezana s članstvom u sindikatu: povezanost doživljaja nesigurnosti posla i članstva u sindikatu bila je pozitivna kod privremenih radnika, ali ne i kod stalnih radnika. Ovakav rezultat može potaknuti sindikate da se ubuduće usmjeru prema strategijama privlačenja privremenih radnika. U sljedećim istraživanjima hipoteze bi se mogle testirati i longitudinalno.

Ključne riječi: povremeni rad, rad na određeno vrijeme, nesigurnost posla, socijalna zaštita, sindikat